

## CS49/249 (Randomized Algorithms), Spring 2021 : Lecture 5

Topic: Median of Means Proof, More General Chernoff Bounds

Disclaimer: These notes have not gone through scrutiny and in all probability contain errors.

Please discuss in Piazza/email errors to deeparnab@dartmouth.edu

- We will complete the proof of the following theorem using the Chernoff bound.

**Theorem 1.** [Boosting Theorem or the Median-of-Means Theorem.]

Let  $\widehat{\text{est}}$  be an *unbiased* estimator of some statistic  $\text{stat}$ . Then, one can obtain an  $(\varepsilon, \delta)$ -multiplicative estimate of  $\text{stat}$  using  $K$  independent samples of  $\widehat{\text{est}}$ , where

$$K = \frac{C \text{Var}[\widehat{\text{est}}]}{(\text{Exp}[\widehat{\text{est}}])^2} \cdot \frac{1}{\varepsilon^2} \cdot \ln\left(\frac{2}{\delta}\right)$$

where  $C$  is some constant. Consequently, one can obtain an  $(\varepsilon, \delta)$ -additive estimate of  $\text{stat}$  using  $K'$  independent samples of  $\widehat{\text{est}}$ , where  $K' = \frac{C \text{Var}[\widehat{\text{est}}]}{\varepsilon^2} \cdot \ln\left(\frac{2}{\delta}\right)$

Recall, last lecture from an unbiased estimator  $\widehat{\text{est}}$ , we obtained an estimator  $\text{est}'$  which was  $(\varepsilon, \frac{1}{3})$ -approximate. That is,

$$\Pr[\text{est}' \notin (1 \pm \varepsilon)\text{stat}] \leq \frac{1}{3} \quad (1)$$

To obtain  $\text{est}'$ , we used  $s \geq \frac{3 \text{Var}[\widehat{\text{est}}]}{\varepsilon^2 (\text{Exp}[\widehat{\text{est}}])^2}$  independent samples of  $\widehat{\text{est}}$ .

- **Boosting using the median.** To obtain the better estimate, we take a bunch of samples from  $\text{est}'$ , and return the *median* of these estimates.

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1: procedure MEDIAN-OF-MEANS: ▷ Assumes access to an unbiased estimate  $\widehat{\text{est}}$ .
2:   for  $i = 1$  to  $c$  independently do: ▷  $c$  is an integer  $\geq 36 \ln(2/\delta)$ .
3:     Run BETTER-ESTIMATOR( $s$ ) with  $s \geq \frac{3 \text{Var}[\widehat{\text{est}}]}{\varepsilon^2 (\text{Exp}[\widehat{\text{est}}])^2}$  to obtain  $\text{est}'_i$ .
4:   Return  $\text{est} \leftarrow$  median of  $(\text{est}'_1, \dots, \text{est}'_c)$ .
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**Observation 1.** The total number of independent samples of  $\widehat{\text{est}}$  used is  $c \cdot s \geq \frac{108 \text{Var}[\widehat{\text{est}}] \ln(2/\delta)}{\varepsilon^2 (\text{Exp}[\widehat{\text{est}}])^2}$

We claim that  $\text{est}$  is an  $(\varepsilon, \delta)$ -estimate. To prove this, let us define some “bad events”. Let  $X_i$  be the *indicator* random variable of the event that the  $i$ th estimate  $\text{est}'_i$  is an over-estimate. That is,

$$X_i = \begin{cases} 1 & \text{if } \text{est}'_i \geq (1 + \varepsilon)\text{stat} \\ 0 & \text{otherwise} \end{cases}$$

Note that  $X_i$ ’s are mutually independent because all the estimates are independent. (1) also gives us  $\text{Exp}[X_i] \leq \frac{1}{3}$ . Define  $X = \sum_{i=1}^c X_i$ , and thus  $\text{Exp}[X] \leq \frac{c}{3}$ . Next comes the main observation.

**Claim 1.** If  $X < \frac{c}{2}$ , then  $\text{est}$  is *not* an over-estimate, implying,  $\Pr[\text{est} \geq (1 + \varepsilon)\text{stat}] \leq \Pr[X \geq \frac{c}{2}]$ .

*Proof.* If less than half of the  $\text{est}'_i$ 's are over-estimates, then the median is *not* going to be an over-estimate.  $\square$

Now the Chernoff bound gives us (using the upper bound<sup>1</sup>  $\frac{c}{3} \geq \mathbf{Exp}[X]$ )

$$\Pr\left[X \geq \frac{(1 + \varepsilon)c}{3}\right] \leq e^{-\frac{\varepsilon^2 c}{9}} \xRightarrow{\varepsilon = \frac{1}{2}} \Pr\left[X \geq \frac{c}{2}\right] \leq e^{-\frac{c}{36}} \leq \underbrace{\delta}_{\text{since } c \geq 36 \ln(2/\delta)} \frac{\delta}{2}$$

Thus, we get  $\Pr[\text{est} \geq (1 + \varepsilon)\text{stat}] \leq \frac{\delta}{2}$ . A similar argument (with “underestimate bad-events”) gives  $\Pr[\text{est} < (1 - \varepsilon)\text{stat}] \leq \frac{\delta}{2}$ . Therefore,  $\Pr[\text{est} \notin (1 \pm \varepsilon)\text{stat}] \leq \delta$ . Which completes the proof.

**Exercise:** Do this argument.

- **The General Chernoff Bound.** The Chernoff bound statement stated last lecture is a corollary to the general statement which looks a little complicated, but over time you just get used to it. Here it is.

**Theorem 2** (General Chernoff: the upper tail). Let  $X_1, X_2, \dots, X_n$  be **independent Bernoulli** random variables with each  $X_i \in \{0, 1\}$ . Let  $X = \sum_{i=1}^n X_i$ . Then, for any  $t > 0$ , we have

$$\Pr[X \geq (1 + t) \mathbf{Exp}[X]] \leq e^{-\mathbf{Exp}[X] \cdot g(t)} \quad (2)$$

where  $g(t) := (1 + t) \ln(1 + t) - t$ .

When  $t \in (0, 1)$ , then one can **show** that  $g(t) \geq \frac{t^2}{3}$ , which in turn implies (UT1) from the last lecture. Similarly, when  $t \in [1, 4]$ , then one can **show** that  $g(t) \geq \frac{t^2}{4}$  implying (UT2), and when  $t > 4$ , then one can **show** that  $g(t) \geq \frac{t \ln t}{2}$  implying (UT3). All this can also be proved using calculus as well. There is a similar statement for the lower tail.

**Theorem 3** (General Chernoff: the lower tail). Let  $X_1, X_2, \dots, X_n$  be **independent Bernoulli** random variables with each  $X_i \in \{0, 1\}$ . Let  $X = \sum_{i=1}^n X_i$ . Then, for any  $0 < t < 1$ , we have

$$\Pr[X \leq (1 - t) \mathbf{Exp}[X]] \leq e^{-\mathbf{Exp}[X] \cdot g(t)} \quad (3)$$

where  $g(t) := (1 - t) \ln(1 - t) + t$ .

One obtains (LT) by **showing**  $g(t) \geq \frac{t^2}{2}$

- **Proof Strategy for Theorem 2 and Theorem 3.** The exact proof of the theorems are not as important as the three main ideas behind it. I would recommend you all to understand these ideas and try to fill the details (which requires some calculussing) yourself. If you fail, then refer to any of the bazillion proofs either in the textbook or on the web. However, this is not the only way to prove these: see [here](#) for five different ways.

<sup>1</sup>see remark following the statement in the previous notes

- a. (Idea 1). Observe that for any parameter  $\gamma > 0$ , the event  $\{X \geq (1+t) \mathbf{Exp}[X]\}$  is the same as  $\{e^{\gamma X} \geq e^{\gamma(1+t) \mathbf{Exp}[X]}\}$ . What does this gain us? Think of  $e^{\gamma X}$  as a random variable  $Z$ . Then  $Z$  is *always* positive, and we can apply Markov on it. This will hold *even when*  $X$  could have been negative (which doesn't happen if all  $X_i$ 's are  $\{0, 1\}$ , but this allows us to prove even more general Chernoff bounds.)
- b. (Idea 2). What does Markov give? If we let  $\theta := e^{\gamma(1+t) \mathbf{Exp}[X]}$  (which note is a fixed quantity with nothing random in it), then

$$\Pr[Z \geq \theta] \leq \frac{\mathbf{Exp}[Z]}{\theta} \quad (4)$$

What is  $\mathbf{Exp}[Z]$ ? Here we use

$$\mathbf{Exp}[Z] = \mathbf{Exp}[e^{\gamma X}] = \mathbf{Exp}[e^{\gamma \sum_{i=1}^n X_i}] = \mathbf{Exp}\left[\prod_{i=1}^n e^{\gamma X_i}\right]$$

and now the kicker, by independence (which implies  $\mathbf{Exp}[f(X_1)f(X_2)] = \mathbf{Exp}[f(X_1)] \mathbf{Exp}[f(X_2)]$ ),

$$\mathbf{Exp}[Z] = \prod_{i=1}^n \mathbf{Exp}[e^{\gamma X_i}] \quad (5)$$

- c. (Idea 3). Now, we use that  $X_i$  is a Bernoulli random variable with some probability  $p_i$  of being 1 and  $(1 - p_i)$  probability of being 0. And therefore,

$$\mathbf{Exp}[e^{\gamma X_i}] = p_i e^{\gamma} - (1 - p_i) = 1 + p_i(e^{\gamma} - 1)$$

And finally we use one bit of analytic trick which is again a pervasive one: we use the inequality  $1 + z \leq e^z$  for any  $z$ , to get

$$\mathbf{Exp}[e^{\gamma X_i}] \leq e^{p_i(e^{\gamma} - 1)} \quad (6)$$

Now one plugs (6) into (5) to get an upper bound on  $\mathbf{Exp}[Z]$ , which is plugged into (4) giving an upper bound on the probability as a function of  $\gamma$ . Then, one uses calculus to find the  $\gamma$  which minimizes this quantity. And after some sweat, one proves [Theorem 2](#). A similar strategy works for [Theorem 3](#). I once again recommend every one doing this exercise from this point on; it is something one needs to do at least once in their lives!

As mentioned above, the similar strategy can actually prove Chernoff style bounds even when the  $X_i$ 's are not  $\{0, 1\}$  variables. Indeed, the *same* bounds hold even when each  $X_i$  is a random variable whose domain is  $[0, 1]$ . And in the proof strategy above, a little more care (analysis) is needed in (Idea 3) to figure out  $\mathbf{Exp}[e^{\gamma X_i}]$ ; in fact one shows the Bernoulli case is the “worst” (which, if you think about it, makes sense as that has the maximum variance).

### Learning Tidbits:

- Algorithm Design: Taking “median” of independent averages/means boosts confidence. This is a general principle.
- Analysis: Seeing how Chernoff bounds apply : express random variable of interest as sum of independent random variables, and then it lingers “very close” to the expectation.

